

The Role of Assumptions in Latent Variable Model Estimation: Elucidation and Tuning

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Introduction: Statistical Problem

- Observed variables ($i=1,\dots,n$): Y_i =M-variate; x_i =P-variate
- Focus: response (Y) distribution = $G_{Y|x}(y|x)$; x-dependence
- Modeling issue: flexible or theory-based?
 - Flexible: $g_m(E[Y_{im}|x_i]) = f_m(x_i)$, $m=1,\dots,M$
 - Theory-based:
 - > Y_i generated from latent (underlying) U_i :
$$F_{Y|U,x}(y|U=u,x;\pi) \quad (\textit{Measurement})$$
 - > Focus on distribution, regression re U_i :
$$F_{U|x}(u|x;\beta) \quad (\textit{Structural})$$
 - > Overall, hierarchical, model:
$$F_{Y|x}(y|x) = \int F_{Y|U,x}(y|U=u,x) dF_{U|x}(u|x)$$

Motivation

The Debate over Mixture and Latent Variable Models

- **In favor:** they
 - acknowledge **measurement problems:** errors, differential reporting
 - **summarize** multiple measures **parsimoniously**
 - operationalize **theory**
 - describe population **heterogeneity**

- **Against:** their
 - **modeling assumptions** may determine scientific conclusions

 - **interpretation** may be ambiguous
 - > nature of latent variables?
 - > comparable fit of very different models
 - > seeing is believing

Possible Approaches to the Debate

- Argue advantages of favorite method
- Hybrid approaches:
 - Parallel analyses (e.g. *Bandeen-Roche et al. AJE 1999*)
 - Marginal mean + LV-based association
(e.g. *Heagerty, Biometrics, 2001*)
- Sensitivity analyses
- **“Popperian”**
 - **Pose parsimonious model**
 - **Learn how it fails to describe the world**

Outline

- Modeling and estimation framework
- Specifying the target of estimation
 - *Supposing that the target uniquely exists ...*
 - > Strategy for delineating it
 - > Validity of the strategy
 - *Unique existence of the target*
- Application: Post-traumatic Stress Disorder
- Refocusing: Methodology to counterbalance competing assumptions

Application: Post-traumatic Stress Disorder Ascertainment

- PTSD

- Follows a qualifying traumatic event

- > *This study: personal assault, other personal injury/trauma, trauma to loved one, sudden death of loved one*
= “x”, along with gender

- Criterion endorsement of symptoms related to the event ⇒ diagnosis

- > Binary report on 17 symptoms = “Y”

- A recent study (Chilcoat & Breslau, *Arch Gen Psych*, 1998)

- Telephone interview in metropolitan Detroit

- n=1827 with a qualifying event

- Analytic issues

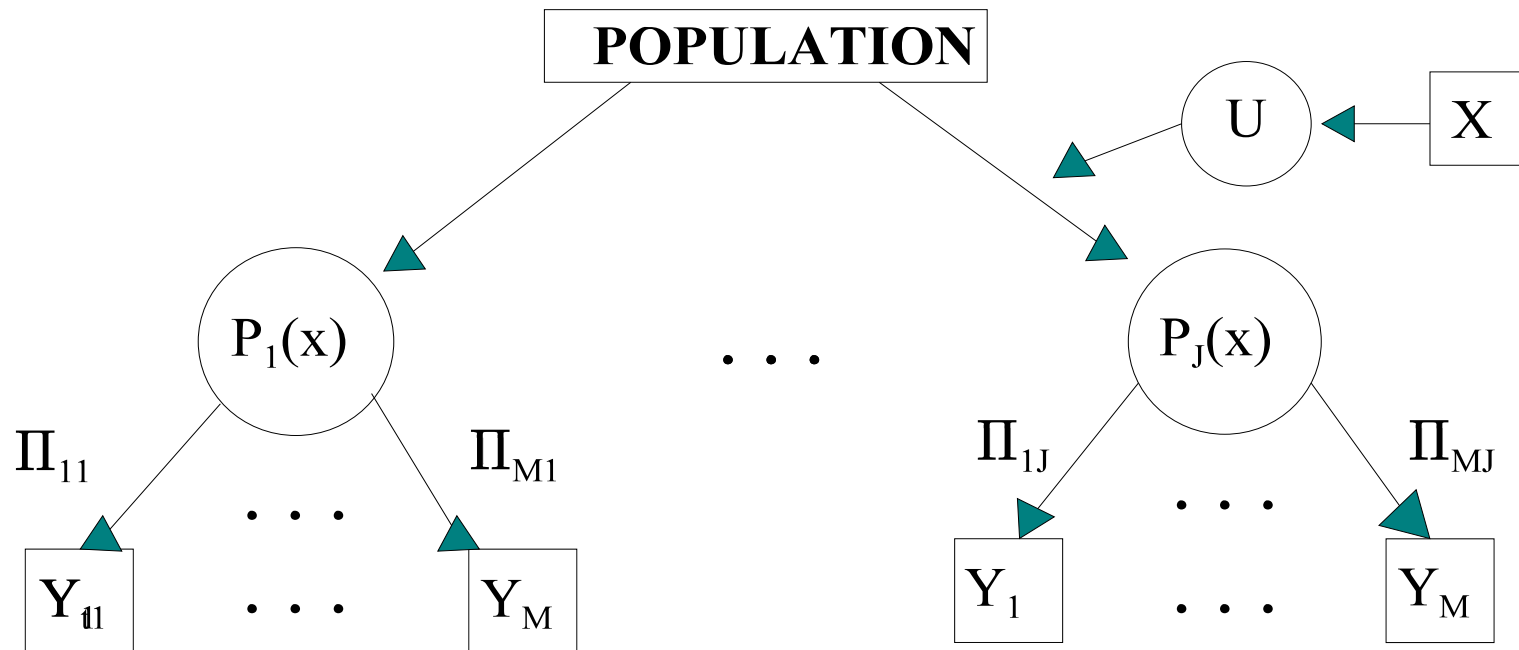
- > Nosology

- > Does diagnosis differ by trauma type or gender?

- > *Are female assault victims particularly at risk?*

Model 1

Latent Class Regression



$$\begin{aligned}
 &> P_j(x) = \Pr\{U = j|x\} \\
 &> \pi_{mj} = \Pr\{Y_m=1|U = j\}
 \end{aligned}$$

References: Dayton & Macready 1988, van der Heidjen et al., 1996; Bandeen-Roche et al., 1997

Latent Class Regression (LCR) Model

- **Model:**

$$f_{Y|x}(y|x) = \sum_{j=1}^J P_j(x, \beta) \prod_{m=1}^M \pi_{mj}^{y_m} (1 - \pi_{mj})^{1-y_m}$$

- **Structural model assumption :** $[U_i|x_i] = \Pr\{U_i=j|x_i\} = P_j(x_i, \beta)$

— $RPR_j = \Pr\{U_i = j|x_i\} / \Pr\{U_i = J|x_i\}; j=1, \dots, J$

- **Measurement assumptions :** $[Y_i|U_i]$

— conditional independence

— nondifferential measurement

> *reporting heterogeneity unrelated to measured, unmeasured characteristics*

- **Fitting:** ML w EM; robust variance (e.g. *Muthén & Muthén 1998, M-Plus*)

- *Posterior* latent outcome info: $\Pr\{U_i=j|Y_i, x_i; \theta=(\pi, \beta)\}$

Methodology

Delineating the Target of Measurement

- **Fit an initial model:** ML, Bayes, etc.
- **Obtain *posterior* latent outcome** info — e.g. $f_{U|Y,x}(u|Y,x;\theta)$
— This talk: empirical Bayes
- **RANDOMLY** generate “empirical LVs,” V_i , according to $f_{U|Y,x}(u|Y,x;\hat{\theta})$
- Analyze V_i AS U_i (accounting for variability in first-stage estimation)
- Estimate measurement structure through empirical analysis of $Y_i|V_i, x_i$

Methodology

Properties “whatever” the True Distribution

- Under Huber (1967)-like conditions:

— Asymptotically:

> Randomization imposes limiting hierarchical model, except
[Y|V,x] arbitrary (and specifiable)

i.e. *underlying variable distribution has an estimable
interpretation even if assumptions are violated*

> No bias in substituting V_i for U_i .

i.e. *regression of V_i on x_i and model-based LV regression
eventually equivalent*

Methodology

More formal statement

- Under Huber (1967)-like conditions:
 - $(\hat{\beta}, \hat{\pi})$ converge in probability to limits (β^*, π^*) .
 - Y_i asymptotically equivalent in distribution to Y^* , generated as:
 - i) Generate U_i^* — distribution determined by (β^*, π^*) , $G_{Y|x}(y|x)$;
 - ii) Generate Y^* — distribution determined by (β^*, π^*) , $G_{Y|x}(y|x)$, U_i^*
 - $\{\Pr[Y_i \leq y | V_i, x_i], i=1,2,\dots\}$ converges in distribution to $\{\Pr[Y_i^* \leq y | U_i^*, x_i], i=1,2,\dots\}$, for each supported y .
 - V_i converges in distribution to U_i^* .

PTSD Study: Descriptive Statistics

Gender	Trauma Type: percentage distribution				n
	<i>Personal Assault</i>	<i>Other Injury</i>	<i>Trauma to loved one</i>	<i>Sudden death</i>	
<i>Male</i>	14.2	37.7	26.9	21.3	964
<i>Female</i>	14.3	26.3	32.2	27.2	863
Total	14.2	32.3	29.4	24.1	1827

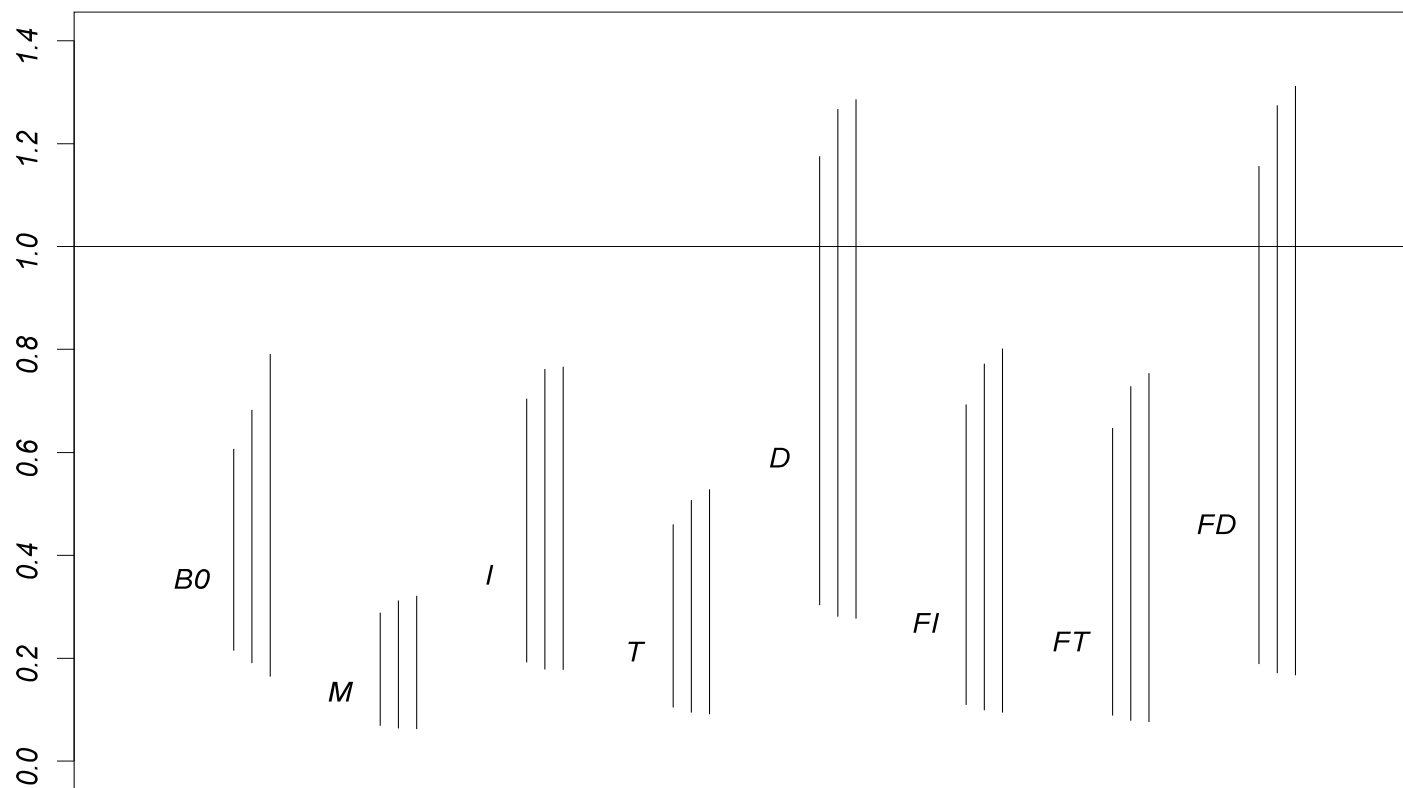
- PTSD symptom criteria met: 11.8% (n=215)
 - By gender: 8.3% of men, 15.6% of women
 - By trauma: *assault (26.9%), sudden death (14.8%), other injury (8.1%), trauma to loved one (6.0%)*
 - Interactions: female x assault (↑), female x other (↓)
 - Criterion issue? 60% reported symptoms short of diagnosis

Latent Class Model for PTSD: 9 items

SYMPTOM CLASS	SYMPTOM (prevalence)	SYMPTOM PROBABILITY (π)		
		Class 1 - NO PTSD	Class 2 - SOME SYMPTOMS	Class 3 - PTSD
RE-EXPERIENCE	Recurrent thoughts (.49)	.20	.74	.96
	Distress to event cues (.42)	.12	.68	.88
	Reactivity to cues (.31)	.05	.51	.77
AVOIDANCE/NUMBING	Avoid related thoughts (.28)	.08	.37	.75
	Avoid activities (.24)	.05	.34	.66
	Detachment (.15)	.01	.14	.64
INCREASED AROUSAL	Difficulty sleeping (.19)	.02	.18	.78
	Irritability (.21)	.02	.22	.83
	Difficulty concentrating (.25)	.03	.30	.89
MEAN PREVALENCE-BASELINE		.52	.33	.14

[Omitted: nightmares, flashback; **amnesia**, ↓ **interest**, ↓ **affect**, **short future**; hypervigilance, startle]

Odds and Relative Odds, with 95% Confidence Intervals

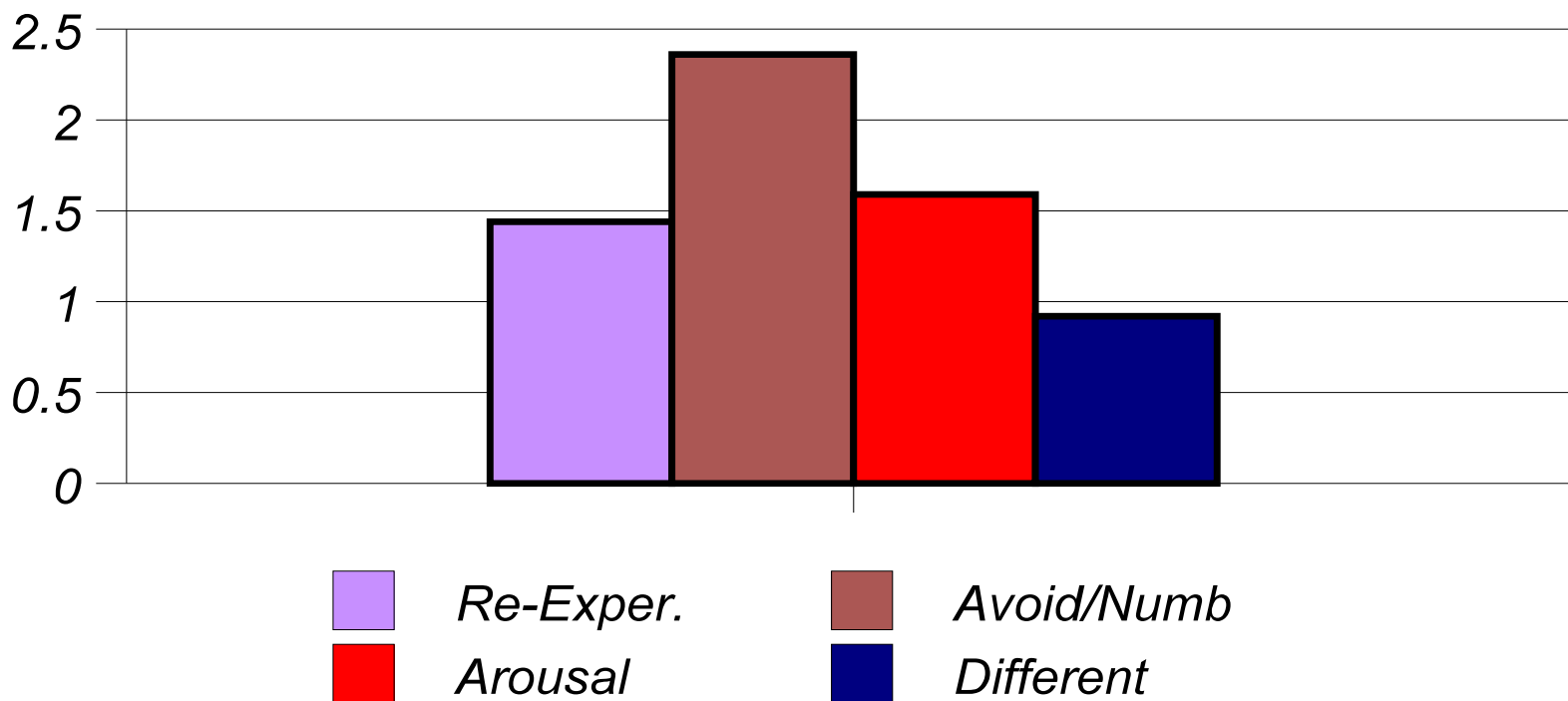


PTSD: DIAGNOSIS, LCR MEASUREMENT MODEL

- Method: Regress item responses on covariates “controlling” for class
— For simplicity: non-assaultive traumas merged into “other trauma”

Variable	Odds Ratio or Interaction Ratio (CI)	By-item Odds Ratio MODEL 2
Female	1.07 (0.93,1.22)	1.07 (0.93,1.22)
Trauma =other than assault (recur.)	3.19 (1.89,5.40)	3.19 (1.89,5.40)
Cue distress x other trauma	0.18 (0.09,0.38)	0.58 (0.36,0.92)
Cue reactivity x other trauma	0.14 (0.07,0.28)	0.44 (0.27,0.72)
Avoid thoughts x other trauma	0.21 (0.11,0.41)	0.68 (0.44,1.05)
Avoid activities x other trauma	0.11 (0.05,0.22)	0.35 (0.21,0.58)
Detachment x other trauma	0.27 (0.13,0.58)	0.88 (0.51,1.49)
Difficulty sleep x other trauma	0.43 (0.21,0.90)	1.37 (0.78,2.42)
Irritability x other trauma	0.28 (0.13,0.61)	0.91 (0.52,1.59)
Concentration x other trauma	0.73 (0.36,1.47)	2.33 (1.35,4.03)

Diagnosis: Conditional Independence (Pairwise ORs)



Re-Exper.	1.44
Avoid/Numb	2.36
Arousal	1.59
Different	0.92

Summary

PTSD Analysis

- The analysis hypothesizes that PTSD is
 - a syndrome comprising **unaffected**, **subclinically affected**, and **diseased** subpopulations of those suffering traumas
 - reported homogeneously within subpopulations
- The hypotheses are consistent with current diagnostic criteria
- Gender x type interactions: are strongly indicated
 - Female assault victims at particular risk
 - ... given the subpopulations defined by the model

Summary

PTSD Analysis

- Symptoms appeared differentially sensitive to different traumas

Within classes: those who had a non-assaultive trauma were

— **less prone** to report distress to cues, reactivity to cues, avoiding thoughts, & avoiding activities

— **more prone** to report recurrent thoughts & difficulty concentrating

- Concern: Current criteria may better detect psychiatric sequelae to assault than to traumas other than assault

Summary

- What I delineated
 - A philosophy
 - > Fit an ideal model
 - > Determine the nature of measurement achieved in fact
 - Theory: On the nature of measurement
 - Methodology: To implement the philosophy
 - New work: On compromise between potentially competing validation criteria
- Strengths / benefits
 - Improved use / usefulness of latent variable models
 - Attention to estimability
 - Allows some distrust of the data

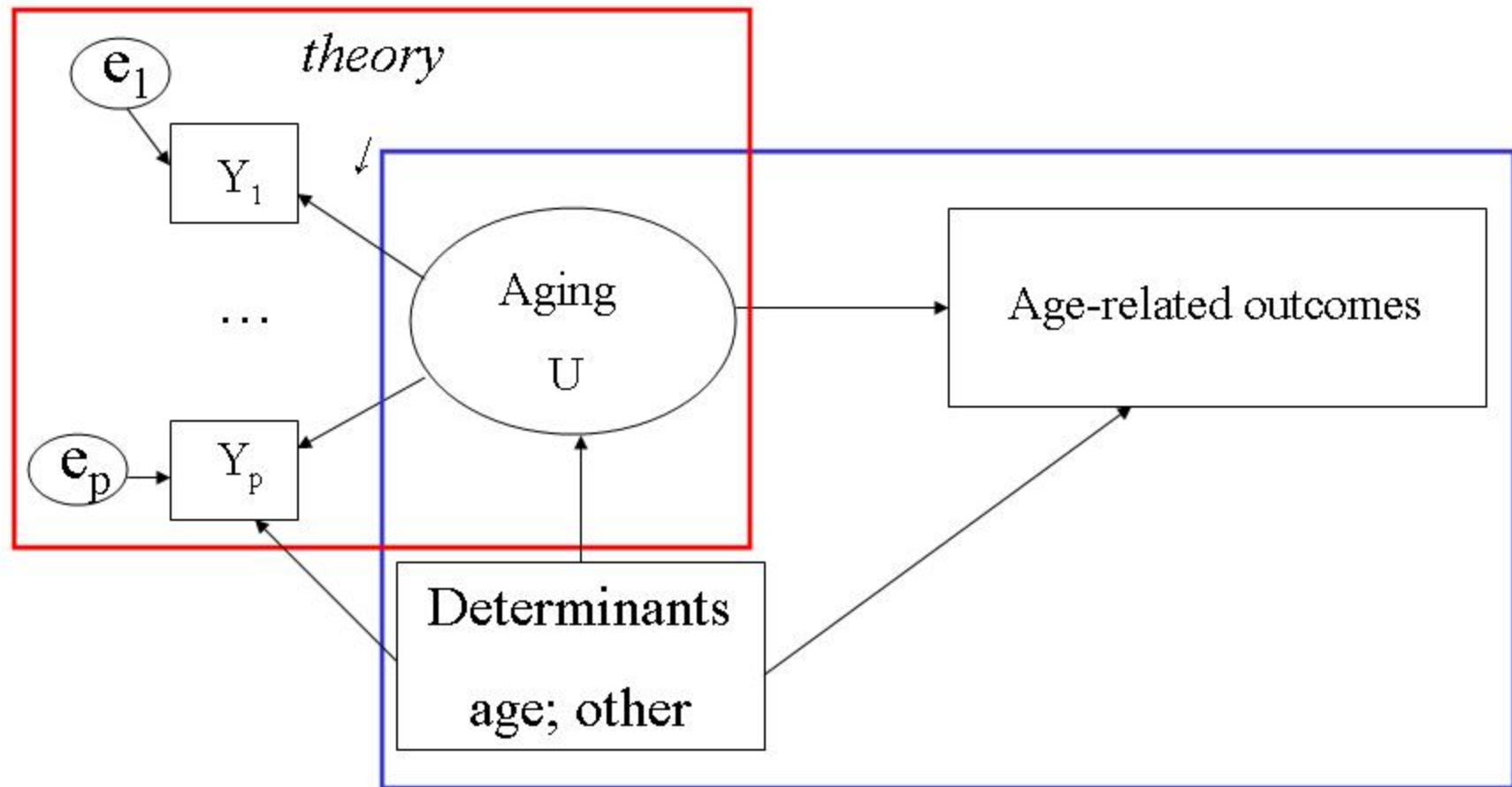
Refocusing of the aim:

Measurement

- A well defined target; a less-well-defined operationalization
- We seek measures that have **validity** for representing their targets
 - LV assumptions encode validity criteria,...
 - ... some better than others
- Objective: Method to unify multiple validation aspects in 1 analysis
 - Balancing potentially conflicting aspects
 - **Today's focus**: “scale” weighing the balance

Measurement

Latent Variable Paradigm



Rationale of the New Work

- Which deserves pre-eminence?
 - Internally validating assumptions?
 - Externally validating assumptions?
 - Some compromise?
- Proposal: Allow compromise via “penalized” fitting

Implementing penalization

- On LCR kernel: Houseman, Coull & Betensky, *BMCS*, 2006
- On LCR mixing distribution: Sheppard Ph.D. thesis
- Key questions
 - Form of the penalty
 - Different purpose than usual?
 - What is the objective function?

Penalization

Very brief background

- Fitting: minimize

$$-2 \ln L(\theta; Y, x) + \lambda g(\theta)$$

- Examples

- “Ridge”: $g(\theta) = \sum_j \theta_j^2$
- “Lasso”: $g(\theta) = \sum_j |\theta_j|$

Green, Int Stat Rev, 1987; Tibshirani, JRSS-B, 1996

Penalization

Very brief background

- A useful equivalence: penalized fit obtains via **formulating parameters as crossed random effects**
 - “Ridge”: $\theta_j \sim N(0, \sigma\{\lambda\}^2)$
 - “Lasso”: $\theta_j \sim \text{double exp}(0, h\{\lambda\})$

Wahba, JRSS-B, 1978; Ngo & Wand, J Stat Software, 2004

Form of the penalty

Current case

- Usual purpose: regularization
- Here: secondary validation
- Discriminant hypothesis:
Genotypes predispose individuals to only one “subtype” of depression

Form of the penalty

Genetic subtypes example

- Say, LCR with one normal class (1) and two disordered classes (2, 3):
- Hypothesis: β_{1j} negligible, and $\beta_{1j'}$ appreciable, in

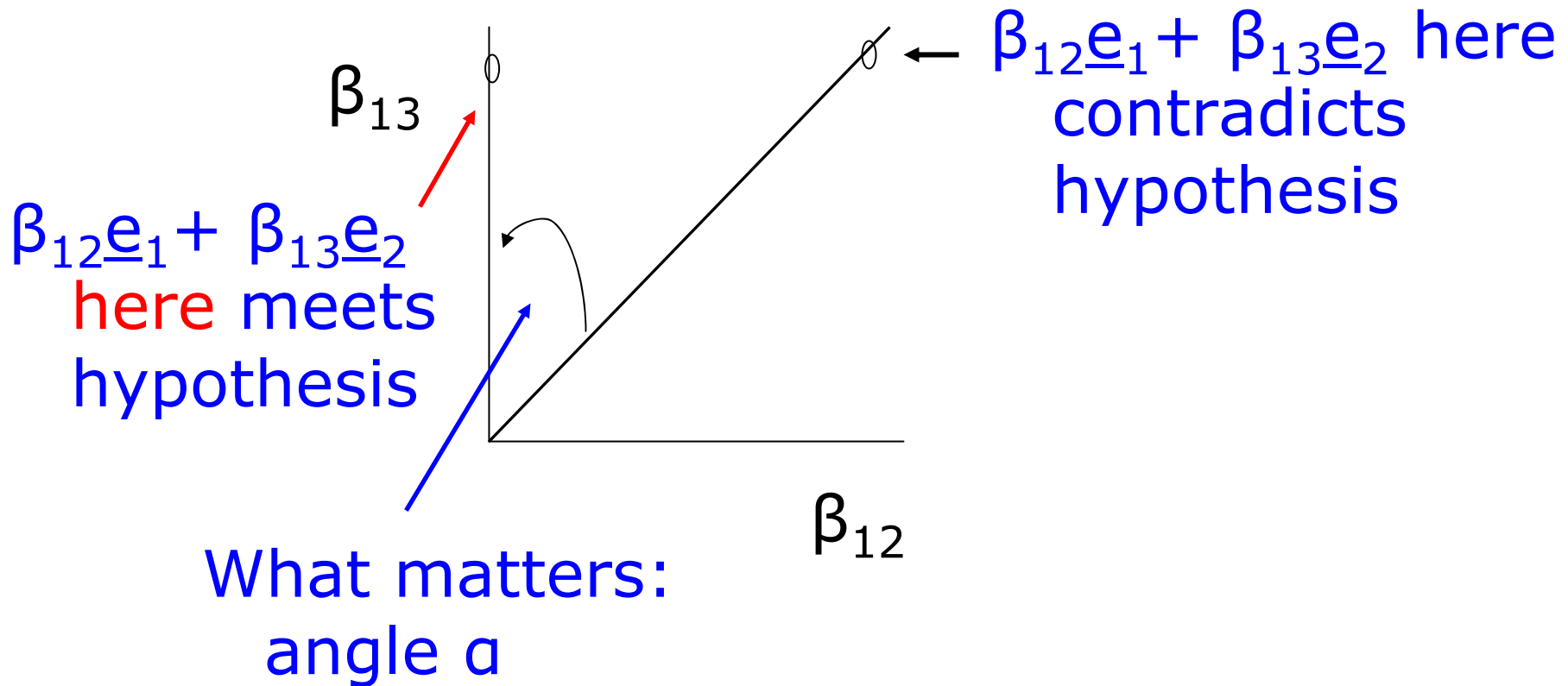
$$\log\left(\frac{p_k}{p_1}\right) = \beta_{0k} + \beta_{1k}x$$

with $p_k = \text{pr}(\text{class } k)$; $x = \text{genotype indicator}$;
 $k=2,3$; $j, j' \in \{2,3\}$; $j \neq j'$

Form of the penalty

Genetic subtypes example

- Ridge, lasso not quite right



Form of the penalty

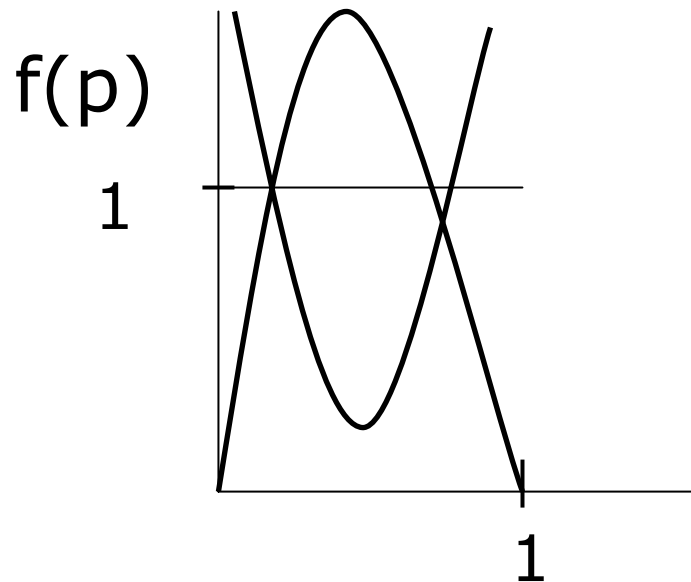
Genetic subtypes example

- Approach 1
 - Consider $\alpha \in [0, 90]$ degrees
 - Desired orientations are $\cos(\alpha)=1$, $\sin(\alpha)=1$
 - i.e., goal: minimize $\cos(\alpha)+\sin(\alpha)$
 - i.e. minimize
$$\frac{|\beta_{12}|+|\beta_{13}|}{\sqrt{\beta_{12}^2 + \beta_{13}^2}}$$

Form of the penalty

Genetic subtypes example

- Approach 2
 - Write $\beta_{12} = p\beta$; $\beta_{13} = (1-p)\beta$
 - Fit with beta random effect on p



- Generalization: $\underline{\beta} = \underline{p}\beta$, $\underline{p} \sim \text{Dirichlet}$

Fitting

Approach 2

- E-M algorithm: quite straightforward
- E-step: Computes posterior class membership probabilities given current parameter iterates
- M-step: minimize (e.g. Nelder-Mead)

$$- \sum_{i=1}^n \sum_{j=1}^J h(j|data) \ln[f_{U|x}(u|x, p, \beta)] + \left(1 - \frac{\Delta}{2}\right) \ln[p(1-p)]$$

Simulation study

Three-class model

- 100 reps; single $x \sim \text{Unif}(-.5, .5)$; $n=1000$
- Poly Log Reg: $\beta_{02}=\beta_{03}=0$; $\beta_{12}=-1.4$;
 $\beta_{13} \in \{0, -1.4\}$
- Measurement:

Class 1	Class 2	Class 3
.15	.85	.85
.15	.85	.85
.15	.85	.85
.15	.13	.85
.15	.13	.85

Simulation study

Three-class model

- Two assumption scenarios
 - Frank LCR
 - Differential measurement: First three items have increased $\log(\text{odds} = 1)$ per unit x of 1.4 in class 3

Simulation Study

Diff. Meas.; $\beta_{13}=0$; $\beta_{12}=-1.4$

Param.	Penalized		LCR	
	Estimate	SE	Estimate	SE
β_{13}	-0.04	0.14	-0.54	0.31
β_{12}	-0.79	0.30	-1.01	0.34

Simulation Study

Non-diff meas; $\beta_{13}=0$; $\beta_{12}=-1.4$

Param.	Penalized		LCR	
	Estimate	SE	Estimate	SE
β_{13}	0	0	0.04	0.32
β_{12}	-1.42	0.35	-1.41	0.38

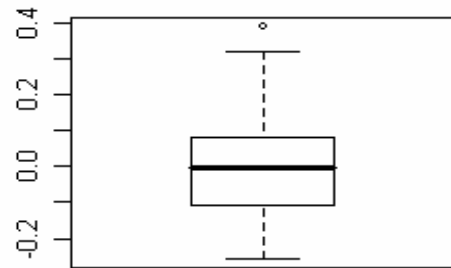
Simulation Study

Non-diff meas; $\beta_{12}=\beta_{13}=-1.4$

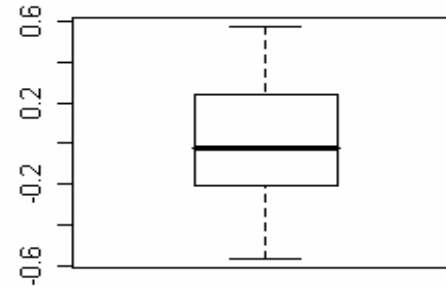
Param.	Penalized		LCR	
	Estimate	SE	Estimate	SE
β_{13}	-1.45	0.34	-1.45	0.30
β_{12}	-1.38	0.31	-1.38	0.31

Simulation Study

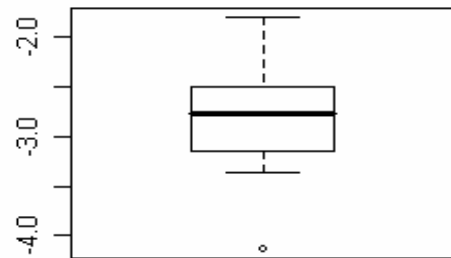
Non-diff meas; $\beta_{12}=\beta_{13}=-1.4$



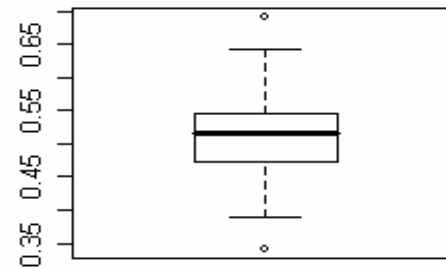
beta01



beta02



beta



p1

Discussion

- A primary issue: Why a hierarchical model at all?
 - PTSD: Why not DSM Y, delineate measurement properties?
- 1) **Nosology**
 - a. Central role of cond. independence, non-diff. measurement.
 - b. Guidance in creating, say, three rather than two groups.
- 2) **The quest for the “ideal”**
 - a. Could have turned out that LCR much less subject to NDM, than DSM: i.e. issue with diagnostic criteria rather than items.
 - b. In fact: LCR and DSM about equally subject to NDM
 - c. Ultimate recommendation: DSM

Discussion

- Beyond delineation of assumptions....
- Further Work: Uniqueness of target
 - Delineation of plausible models
 - Displays, complicated models
 - ***Implication***: Guidance on parsimony versus complexity
- Further work: Big picture for validation compromise
 - How does measurement conform?
 - How should one determine the magnitude of the compromise?
 - Empirical adjudication
 - Clinical / scientific utility
 - Ultimate: Gold standard (aging: telomeric shortening)
- Why not be Bayesian?

Implications

- Refined understanding of health states and their measurement
 - Integrating biology
 - Increasing sensitivity, specificity
- Heightened accuracy, precision for
 - Delineating etiology
 - Developing and targeting interventions

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